

## Expected Value Theorems

A random variable is a variable whose value is a numerical outcome of a random process.

**Definition 1:** The expected value of the discrete random variable  $X$ , denoted  $E(X)$ , is defined to be  $E(X) = \sum_i x_i p(x_i)$ . This is another way to define the mean of the population, so  $E(X) = \mu$ .

**Definition 2:** The variance of the discrete random variable  $X$ , denoted  $V(X)$ , is defined to be  $V(X) = E((x - \mu)^2)$ .

### Expected Values of Linear Transformations

**Theorem 1:**  $E(aX) = aE(X)$

**Proof:**  $E(aX) = \sum_i ax_i p(x_i) = a \sum_i x_i p(x_i) = aE(X)$ .

**Theorem 2:**  $E(X + b) = E(X) + b$

**Proof:**  $E(X + b) = \sum_i (x_i + b)p(x_i) = \sum_i x_i p(x_i) + \sum_i bp(x_i)$   
 $= \sum_i x_i p(x_i) + b \sum_i p(x_i) = E(X) + b$ , since  $\sum_i p(x_i) = 1$ .

### Expected Values of Sums and Products

**Theorem 3:**  $E(X + Y) = E(X) + E(Y)$

**Proof:**  $E(X + Y) = \sum_i \sum_j (x_i + y_j) p(x_i) p(y_j)$   
 $= \sum_i \sum_j x_i p(x_i) p(y_j) + \sum_i \sum_j y_j p(x_i) p(y_j)$

but  $x_i p(x_i)$  is constant in the summation over  $j$ , and  $y_j p(y_j)$  is a constant in the summation over  $i$ . So we have

$$\sum_i x_i p(x_i) \sum_j p(y_j) + \sum_i p(x_i) \sum_j y_j p(y_j) = E(X) \cdot 1 + 1 \cdot E(Y) = E(X) + E(Y).$$

(Note that this theorem does not require the random variables  $X$  and  $Y$  to be independent.)

**Theorem 4:**  $E(XY) = E(X)E(Y)$ , when  $X$  and  $Y$  are independent random variables.

**Proof:** Same as Theorem 3.

### Linear Transformations and Variance

**Theorem 5:**  $V(X) = E(X^2) - (E(X))^2$

**Proof:** 
$$V(X) = E((X - \mu)^2) = E(X^2 - 2X\mu + \mu^2) = E(X^2) - 2\mu E(X) + E(\mu^2)$$
$$= E(X^2) - 2\mu^2 + \mu^2 = E(X^2) - \mu^2 = E(X^2) - (E(X))^2.$$

**Theorem 6:**  $V(aX) = a^2V(X)$

**Proof:** 
$$V(aX) = E(a^2X^2) - (E(aX))^2 = a^2E(X^2) - (aE(X))^2$$
$$= a^2E(X^2) - a^2(E(X))^2 = a^2V(X).$$

**Theorem 7:**  $V(X + b) = V(X)$

**Proof:** 
$$V(X + b) = E(X + b)^2 - (E(X + b))^2 = E(X^2 + 2Xb + b^2) - (E(X) + b)^2$$
$$= E(X^2 + 2Xb + b^2) - [(E(X))^2 + 2bE(X) + b^2] = V(X).$$

### Variance of Sums and Differences

**Theorem 8:**  $V(X + Y) = V(X) + V(Y)$ , when the random variables  $X$  and  $Y$  are independent.

**Proof:** 
$$V(X + Y) = E(X + Y)^2 - [E(X + Y)]^2$$
$$= E(X^2 + 2XY + Y^2) - [(E(X))^2 + 2E(X)E(Y) + (E(Y))^2]$$
$$= E(X^2) + E(Y^2) - (E(X))^2 - (E(Y))^2 = V(X) + V(Y).$$

So,  $\sigma_{X+Y} = \sqrt{\sigma_X^2 + \sigma_Y^2}$ .

**Theorem 9:**  $V(X - Y) = V(X) + V(Y)$ , when the random variables  $X$  and  $Y$  are independent..

**Proof:** Same as Theorem 8.

Suppose we generate  $n$  independent values of a random variable  $X$ . What is the expected value and variance of the sum?

**Theorem 10:**  $E(X_1 + X_2 + X_3 + \cdots + X_n) = nE(X)$

**Proof:**  $E(X_1 + X_2 + \cdots + X_n) = E(X_1) + E(X_2 + \cdots + X_n)$   
 $= E(X_1) + E(X_2) + E(X_3 + \cdots + X_n) = E(X_1) + E(X_2) + \cdots + E(X_n) = nE(X)$ .

**Theorem 11:**  $V(X_1 + X_2 + X_3 + \cdots + X_n) = nV(X)$  (recall that all  $X_i$  are independent)

**Proof:**  $V(X_1 + X_2 + \cdots + X_n) = V(X_1) + V(X_2 + \cdots + X_n)$   
 $= V(X_1) + V(X_2) + V(X_3 + \cdots + X_n) = V(X_1) + V(X_2) + \cdots + V(X_n) = nV(X)$ .

So,  $\sigma_{X_1+X_2+\cdots+X_n} = \sigma_X \sqrt{n}$ .

### Sampling Distribution of the Mean

What is the mean and standard deviation of the sampling distribution of the mean?

**Theorem 12:**  $E(\bar{X}) = \mu$

**Proof:**  $E(\bar{X}) = E\left(\frac{X_1 + X_2 + \cdots + X_n}{n}\right) = \frac{1}{n}E(X_1 + X_2 + \cdots + X_n) = \left(\frac{1}{n}\right)(nE(X)) = \mu$ .

**Theorem 13:**  $V(\bar{X}) = \frac{\sigma_X^2}{n}$

**Proof:**  $V(\bar{X}) = V\left(\frac{X_1 + X_2 + \cdots + X_n}{n}\right) = \frac{1}{n^2}V(X_1 + X_2 + \cdots + X_n) = \left(\frac{1}{n^2}\right)(nV(X)) = \frac{\sigma_X^2}{n}$ .

So  $\sigma_{\bar{X}} = \frac{\sigma_X}{\sqrt{n}}$ .

## Sample Variance

Why is the denominator of the sample variance  $n-1$ ? We use  $n-1$  because the value of

$\frac{\sum_i (X_i - \bar{X})^2}{n-1}$  is an unbiased estimator of the population variance  $\sigma_X^2$ .

**Theorem 14:**  $E\left(\frac{\sum_i (X_i - \bar{X})^2}{n-1}\right) = \sigma_X^2$ .

**Proof:** We know that  $\sigma^2 = E(X - \mu)^2$ . Consider  $E\left(\sum_i (X_i - \bar{X})^2\right)$ .

$$\begin{aligned} E\left(\sum_i (X_i - \bar{X})^2\right) &= E\left(\sum_i [(X_i - \mu) + (\mu - \bar{X})]^2\right) \\ &= \sum_i E((X_i - \mu)^2) + 2E\left(\sum_i (X_i - \mu)(\mu - \bar{X})\right) + \sum_i E(\mu - \bar{X})^2 \\ &= \sum_i \sigma_X^2 + 2E\left(\sum_i X_i \mu\right) - 2E\left(\sum_i X_i \bar{X}\right) - 2E\left(\sum_i \mu^2\right) + 2E\left(\sum_i \mu \bar{X}\right) + \sum_i \frac{\sigma_X^2}{n} \\ &= n\sigma_X^2 + 2n\mu^2 - 2nE(\bar{X}^2) - 2n\mu^2 + 2n(E(\bar{X}))^2 + n\frac{\sigma_X^2}{n} \\ &= (n+1)\sigma_X^2 + 2nV(\bar{X}) = (n+1)\sigma_X^2 - 2n\left(\frac{\sigma_X^2}{n}\right) = (n-1)\sigma_X^2. \end{aligned}$$

So, we have  $E\left(\sum_i (X_i - \bar{X})^2\right) = (n-1)\sigma_X^2$ . Consequently,  $E\left(\frac{\sum_i (X_i - \bar{X})^2}{n-1}\right) = \sigma_X^2$  and

$(n-1)$  is exactly the right value to produce an unbiased estimate of the population variance from a sample of  $n$  independent values of a random variable.

### Reference:

Goldberg, Samuel, *Probability: An Introduction*, Dover Publications, New York, 1960.